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A Geographic Economy Model in a
Comparative Study**

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Summary

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Keywords: FDI, Environmental Regulation, Economic Geography, Pollution Haven Hypothesis

JEL Classification: F12, F18, Q28

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The pollution haven hypothesis: a geographic economy model in a comparative study*

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Abstract

Although based on theoretical foundations, the pollution haven hypothesis has never been clearly proven empirically. In this study, we reexamine this hypothesis by a fresh take on both its theoretical and empirical aspects. While applying a geographic economy model on French firm-level data, we confirm the hypothesis for the global sample. Through sensitivity analysis, we validate it for Central and Eastern European countries, emerging and high-income OECD countries, but not for the major part of the Commonwealth of Independent States countries. Finally, we show that the pollution haven hypothesis is confirmed in the strongest manner for emerging economies.

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1 Introduction

There is, by now, quite an extensive literature on the factors that influence firms' location decisions abroad. Among these factors, the most studied are the cost of production factors such as labor and capital, and the market access. Additional factors such as taxation or agglomeration effects have also been studied. More recently, a new factor appeared as a potential determinant of firms' location abroad: environmental regulation. This factor was notably put in evidence by Copeland and Taylor (2004) through a simple model of specialization and trade according to which the rich countries, that protect their environment, should abandon their polluting activities

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to developing countries, whose environmental legislation and enforcement are not severe. This statement illustrates the commonly studied "pollution haven hypothesis" (PHH). Nevertheless, empirical research often fails to prove this hypothesis. Furthermore, an approach based on the classic theory of endowments would yield an opposite conclusion: polluting activities are generally capital-intensive and should thus locate in rich countries where capital is abundant.

The debate on the pollution haven hypothesis produced a political challenge by trying to find clear empirical evidence in order to prove or refute what is really a complex and dynamic issue: how does environmental regulation interact with more and more mobile production?

Generally, statistical studies prove that the pollution haven hypothesis cannot be clearly identified. Four potential problems in this literature require more empirical tests. First of all, most studies lack theoretical foundations for the construction of the equations to be tested, which often entails specification errors. Secondly, the studies of Zhang and Markusen (1999) and Cheng and Kwan (2000) demonstrate the importance of relative endowments of production factors in the explanation of the foreign direct investment (FDI). The absence of this determinant can lead to omitted variable bias. Next, Levinson and Taylor (2008) and Keller and Levinson (2002) emphasize the empirical importance of controlling for the unobservable characteristics of industries and locations. Finally, as noted by Smarzynska and Wei (2004), several studies use very aggregated data on FDI, and proxies of the severity of the environmental policy that are far off the real variable to be taken into account, which generally results in bias induced from measurement-error.

In this study we considered these various limits and tried to remedy them by different means. Consequently, we present a classic theoretical model of geographic economy (open economy model with increasing returns to scale, monopolistic competition and trade costs), which supplies us a testable equation for determinants of the firms' location choice, among which we distinguish the impact of the environmental regulation. According to our knowledge, at the moment, there is no empirical study on the link between FDI and environment, in particular on the pollution haven hypothesis, based on a theoretical model of geographic economy. Such models are used in some purely theoretical studies, without empirical estimation of the forces at work (e.g., Rauscher, 2005; Conrad, 2005; Van Marrewijk, 2005). The geographic economy model presents the advantage to take into account the complexity of FDI determinants, such as: production factor endowments (labor, capital, etc.); distance between trade partners, local market size and access to other important markets (market potential of the host country); cultural, historic, linguistic connections. While representing the demand, the market potential derived from theory has the advantage, in comparison to other often used variables (GDP, population, etc.), to take into account the accessibility to the host country's market and that of its neighbors, the trade barriers and competition. Besides, the additional interest of this model is that it enables us to introduce the environmental regulation as a determinant of the location decision. Thus, for the specification of the location choice, we are inspired by the study of Head and Mayer (2004), to which we introduce environmental regulation. Besides labor and capital, we consider pollution as a production factor, the cost of which is the pollution tax

established exogenously by the government. Furthermore, in our empirical model, the fixed costs corresponding to the launch of an activity / creation of a firm differ across countries and thus have an impact on the firm's location choice. These costs are proxied in our study by the institutional quality of the potential locations.

For the empirical test we use firm-level data, in particular, data on French firms. Our focus is on transition countries, while also considering other groups of countries. In fact, if in 1992 the transition and developing countries counted 30 % of the French multinationals' subsidiaries, in 2002 these countries were the destination of 45% of the French firms abroad. Concerning the specific case of the Central and Eastern European Countries (CEECs) and the countries of the Commonwealth of Independent States (CIS), between 1992 and 2002 the French multinationals multiplied by six the number of their subsidiaries in this region in order to represent 11% of the total of the French firms abroad (3 % in 1992)¹. Regarding emerging countries, while they counted about 25% of the French establishments in 1992, in 2002 they counted about 35%. This important reorientation of the French FDI towards countries with an environmental regulation less severe than that of industrial nations represents an interesting case study for the research on the pollution haven hypothesis.

Furthermore, to represent the environmental regulation's stringency in a more complete way, we create a complex and dynamic index which estimates the relative severity of the environmental policy between countries, based on a diverse set of variables.

Finally, in order to take into account the specific characteristics of industries and countries, the empirical estimations are performed by controlling for different types of industrial sectors (according to their pollution intensity) and groups of countries (transition CEECs, transition countries of CIS, emerging countries², and high-income OECD countries³).

The contribution of this article lies in the empirical estimation and test of the pollution haven hypothesis, using firm level data on worldwide location of French manufacturing subsidiaries. Our empirical results show that looking at diverse countries (pooled sample), the French firms of the manufacturing sector, and in particular the ones belonging to more polluting sectors, prefer to settle in countries where the environmental regulation is the most lenient, thereby confirming the pollution haven hypothesis. Sensitivity analysis including interaction terms between the environmental regulation and the different country groups generate three major conclusions. First of all, we validate this hypothesis for all the CEECs, emerging and high-income OECD countries in our sample. Concerning the CIS transition countries, we discover a prevailing opposite relationship: it is rather the lenient environmental regulation that deters the foreign investments. Finally, we show that emerging economies constitute the group of countries for which this hypothesis is most clearly confirmed.

¹Source : Survey on French subsidiaries implanted abroad, managed by the Directorate of Treasury and Economic Policy on 2003.

²An emerging country is a country, up to there under developed, which undertaken measures and accumulated means, in particular legal and cultural, in order to begin a phase of fast growth of the production and social welfare. According to *Morgan Stanley Emerging Markets Index*, in July 2006, the status of emerging country was awarded to the following countries: Argentina, Brazil, Chile, China, Colombia, Egypt, Iran, India, Indonesia, Israel, South Korea, Malaysia, Mexico, Morocco, Pakistan, Peru, the Philippines, South Africa, Sri Lanka, Taiwan, Thailand, Tunisia, Turkey, Venezuela.

³High-income OECD members except Czech Republic and Republic of Korea, considered here as transition and emerging countries, respectively (see Appendix C for the list of countries included in our sample).

This study is structured as follows: in section 2 we present the relevant literature analyzing the relation between FDI and environmental regulation. In the third section, we describe briefly the theoretical model of the new geographic economy giving the econometric specification for the analysis of the determinants of firms' location decision. The fourth section describes the empirical model and data used. In the fifth section we analyze the empirical results and provide some extended analyses. The last section concludes.

2 Review of the literature

The increasing relocation of industries towards developing countries raises some questions, mainly in the field of employment, and more recently concerning the environment. The relatively less severe environmental regulation in developing countries can create a comparative advantage in pollution-intensive production, and trade openness and FDI could then damage the environment of the host country. From a theoretical point of view, researchers proposed two main hypotheses to explain the direct impact of international trade and FDI on the environment: the "pollution haven hypothesis" and the "factor endowment hypothesis". The first hypothesis assumes that countries are identical except for exogenous differences in their environmental policies. Thus it is cheaper to produce pollution-intensive goods in countries with a weaker environmental regulation, usually poorer countries. Trade, inferred by differences in environmental policy, thus creates pollution havens in the poorer countries. The "factor endowment hypothesis" is the main alternative to the pollution haven hypothesis. It suggests that trade is determined by the relative abundance of production factors (labor and capital in most models) in each country. Thus, if pollution-intensive goods are generally more capital intensive, they should be produced in the developed countries, instead of the developing countries. At the same time, the developed countries are supposed to be abundant in capital and have a stricter environmental policy, unlike poor countries. This illustrates the narrow connection between both hypotheses, which should be taken into account while testing anyone of them.

Unfortunately, the empirical validation of the pollution haven hypothesis is always rather delicate. One of the founder-articles is Grossman and Krueger (1993), which however shows that the trade liberalization between the United States and Mexico in the 1980's and 1990's did not come along with a relocation of polluting industries, which could be explained by the very weak weight of environmental costs imposed on American firms during this period. Since then, articles on this subject followed without a consensus being established, while concerns abound about the effects of environmental standards on trade flows and FDI. However, the scientific research did not manage to prove that environmental regulation affects trade or firms' investment decisions (e.g., Jaffe et al., 1995; Wheeler, 2001).

Ederington et al. (2005) explain partially why previous studies did not confirm the pollution haven hypothesis. They recall that international trade is essentially made between developed countries, whose regulation is quite similar. However, if one examines only the flows between industrial nations and developing countries, the environmental standards have more pronounced effects on the trade structure: with the strengthening of the

environmental regulation of the United States, imports from developing countries decrease. In fact, Ederington et al. (2005) notice that polluting industries are generally the least mobile geographically and thus, it becomes more expensive to establish production in countries that apply a less rigorous regulation.

Most studies in this field use data on trade flows while analyzing the pollution haven hypothesis. Some more recent papers examine this hypothesis by using data on FDI. Eskeland and Harrison (2003) study the effect of the abatement cost and pollution intensity on FDI in Morocco, Cote d'Ivoire, Venezuela and Mexico, and find essentially no empirical support for the pollution haven hypothesis. Besides, they find that the United States factories are appreciably more efficient in terms of energy use and employ more "clean" types of energy than the host country's plants. In a group of 24 transition countries, Smarzynska and Wei (2004) find some, however relatively weak, proof for the pollution haven hypothesis.

Dean et al. (2004), in a study on China, discover a relationship between FDI and environmental regulation completely different from that evoked by the pollution haven hypothesis. In fact, the authors find that a less stringent regulation is a significant determinant for Chinese villages' attractiveness for joint ventures with partners from Hong-Kong, Macao, Taiwan and other countries of South Asia. On the contrary, industrial nations (United States, Japan, United Kingdom etc.) are attracted by higher standards, regardless of industry pollution intensity. The authors suggest that this result could be explained by technological differences.

Other studies, however, assert that environmental regulation influences the spatial allocation of capital. A seminal paper in this literature is that of Keller and Levinson (2002). What distinguished this work was its using of, on the one hand, panel data on inward FDI flows in the United States over a long period of time and, on the other hand, an innovative measure of the relative abatement costs across the States. By applying standard parametric models on panel data, the authors find a robust result showing that abatement costs had moderate dissuasive effects on foreign investments. Given the implications of such a conclusion for trade and environmental policies, any evaluation of the sensibility of these results to changes in the parametric hypotheses is well justified. The application to Keller's and Levinson (2002) data of the recently developed non parametric techniques (Henderson and Millimet, 2007) reveals two results: first, some of the parametric results are not robust, and second, the impact of relative abatement costs is not uniform across the States, and is generally of a smaller magnitude than previously suggested.

Finally, a recent study (Cole et al., 2006) shows the existence of an inverse relationship between FDI and environmental regulation: it is FDI that influences the environmental policy, but this effect is a function of the degree of corruption in the host country. The authors show that at high (low) corruption levels, FDI leads to less (more) stringent environmental policy.

Different theoretical and empirical factors can explain the lack of robust empirical proof for the pollution haven hypothesis. One reason for which developing countries do not tend to become pollution havens can be the fact that the stringency of the environmental standards is only one, and maybe not the most important, factor

determining the comparative advantages between countries. In particular, the endowments of factors such as qualified human capital and physical capital determine largely industrial location and the products that a country will export. As far as the strongly polluting industries tend also to be intensive in capital, the relative lack of capital in the developing countries can prevail over the advantage of low abatement costs (e.g., Antweiler et al., 2001; Cole and Elliott, 2002).

It is thus difficult to conclude on this point, in so far as the laxness of the environmental policy, supposed to attract polluting firms, may be associated with other characteristics that, in their turn, generally discourage the establishment of foreign firms. It is notably the case of weak institutions, expressed through high corruption level, lack of civil freedoms' and property rights' protection, etc. That is why, as Smarzynska and Wei (2004) underline, it is necessary to take into account the effect of institutions on FDI, while trying to test the pollution haven hypothesis.

3 Theoretical background

The theoretical frame of our model is based on the classic hypotheses of the new geographic economy, open economy model with increasing returns to scale, monopolistic competition and trade costs. Monopolistic competition is a common market form, characterized by the following: there are many producers and many consumers in a given market; consumers have clearly defined preferences and sellers attempt to differentiate their products from those of their competitors; goods and services are heterogeneous, usually (though not always) intrinsically so; there are few barriers to entry and exit; producers have a degree of control over price. This frame seems relevant to our study since these criteria characterize well the French market.

Moreover, the existence of intra-industry trade is typically associated with increasing returns to scale, product differentiation and monopolistic competition (Helpman and Krugman, 1985), and the volume of such trade is expected to be higher the greater the equality of trading partners' GDP per capita (Helpman, 1987). Hummels and Levinsohn (1995), however, in repeating Helpman's (1987) analysis for a sample of non-OECD countries, found that gravity models still worked well for a group of countries whose bilateral trade was more likely characterized by homogeneous goods. Evenett and Keller (2002) have argued that the gravity model can nest both the increasing returns/product differentiation story as well as a more conventional homogeneous goods/relative factor abundance story. Thus, we can use the monopolistic competition model in order to explain French multinationals' behavior concerning location in both the developed and developing countries.

In our model, the world consists of $i = 1, \dots, N$ open economies. In every country there are two sectors - industry and agriculture. Given the interest of this paper in the industry's location, we model the second sector as simply as possible. Namely, the traditional sector A is supposed to produce a homogeneous good under Walrasian conditions (constant returns to scale and perfect competition), which is traded costlessly. The manufacturing sector M produces a continuum of differentiated goods, called varieties v , under increasing returns to scale

in an environment of monopolistic competition à la Dixit and Stiglitz (1977). Each firm produces a distinct variety. Let us denote the elasticity of substitution between two varieties $\sigma > 1$. The shipping of these varieties towards another country implies "iceberg" transport costs: $\tau > 1$ units must be sent so that a unity arrives at the destination; the rest, $\tau - 1$, is melt in transit. The consumers spend a part $0 < \mu < 1$ of their income E on the purchase of the composite good M , the rest is spent on the good A , and they have constant elasticity of substitution (CES) sub-utility functions for the composite good. Under constraints of income and each variety's price, the maximization of this sub-utility results in the following demand function of the country's j consumers for a specific variety h produced in a country i (see details in Appendix A):

$$q_{ij}(h) = \frac{[p_i(h) \tau_{ij}]^{-\sigma}}{\sum_i \int_{n_i} [p_i(v) \tau_{ij}]^{1-\sigma} dv} \mu E_j \quad (1)$$

where p_i is the price of the variety h , $h \in [1..n_i]$, in the exporting country i ; n_i is the mass of varieties in any country i , $i \in N$; and τ_{ij} is the trade cost supported by the consumer in the importing country j . The trade cost includes all transaction costs connected to the shipment of goods in space and across borders.

One of the most important factors for the firm's location decision is the production cost. The total production cost in every potential location is supposed to take the following form: $c_i q_i + F_i$. The increasing returns result from fixed costs $F_i(h)$ specific to the firm producing a variety h (further named firm h) and related to launching activity in the country i , $q_i(h)$ is the total production of the firm h in the country i and $c_i(h)$ is its constant production marginal cost. A firm tries to maximize its gross profit on every market. We can write the gross profit realized in any country j by a firm h implanted in a country i :

$$\pi_{ij}(h) = [p_i(h) - c_i(h)] \tau_{ij} q_{ij}(h) \quad (2)$$

In this model à la Dixit and Stiglitz (1977) and Krugman (1980), the production price is a simple mark-up on the marginal cost: $p = c \frac{\sigma}{\sigma-1}$. By substituting it in the equation (1), we obtain the following expression for the quantity that a firm h produces in country i and would ship to any destination j :

$$q_{ij}(h) = \frac{\sigma - 1}{\sigma} \frac{[c_i(h) \tau_{ij}]^{-\sigma}}{\sum_i \int_{n_i} [c_i(v) \tau_{ij}]^{1-\sigma} dv} \mu E_j \quad (3)$$

By replacing the expression (3) and the price expression in the equation (2), by summarizing the gross profits realized by a firm h located in country i while shipping its goods to any market j (equation 2) and by deducting the firm's fixed cost $F_i(h)$, we obtain the total net profit that a firm h could earn in any potential location i :

$$\Pi_i(h) = \frac{c_i(h)^{1-\sigma}}{\sigma} \sum_j \phi_{ij} \frac{\mu E_j}{G_j} - F_i(h) = \frac{c_i(h)^{1-\sigma}}{\sigma} MP_i - F_i(h) \quad (4)$$

with $\phi_{ij} = \tau_{ij}^{1-\sigma}$, $G_j = \sum_i \int_{n_i} [c_i(v) \tau_{ij}]^{1-\sigma} dv$ and $MP_i = \sum_j \phi_{ij} \frac{\mu E_j}{G_j}$. The location choice is thus deter-

mined by the comparison of the characteristics of every potential location: demand (represented by MP_i , an abbreviation of the Market Potential concept developed in section 4), cost of launching an activity/ creating an enterprise, and production cost.

Taking logarithms and denoting $\Lambda_i = \frac{c_i(h)^{1-\sigma}}{\sigma} MP_i$, we write the profitability U of a firm h located in country i :

$$U_i(h) \equiv \ln \Pi_i(h) = \ln(\Lambda_i - F_i(h)) = \ln \Lambda_i + \ln\left(1 - e^{(\ln F_i(h) - \ln \Lambda_i)}\right) \quad (5)$$

We have thus:

$$U_i(h) \equiv \ln MP_i - (\sigma - 1) \ln c_i(h) - \ln \sigma + \ln\left(1 - e^{(\ln F_i(h) - \ln \Lambda_i)}\right) \quad (6)$$

In our model, there are three production factors: K - capital, L - labor and $Poll$ - pollution (the last factor is used only in industry). One of the most common forms used to represent the cost function is that of Cobb-Douglas with constant returns: $c = \frac{1}{A} w^\alpha r^\beta t^\theta$, where $\theta = 1 - (\alpha + \beta)$ and w , r and t are costs of labor, capital and pollution, respectively. The share of labor in the firm's production process is noted by α , of the capital is β and finally that of the pollution is θ , whereas A represents the total factor productivity. With this last specification we can rewrite equation (6) in the following way:

$$U_i(h) \equiv \ln MP_i + (\sigma - 1) \ln A_i - \alpha(h) (\sigma - 1) \ln w_i - \beta(h) (\sigma - 1) \ln r_i - \theta(h) (\sigma - 1) \ln t_i - \ln \sigma + \ln\left(1 - e^{(\ln F_i(h) - \ln \Lambda_i)}\right) \quad (7)$$

The equation (7) predicts that the profitability of a firm settled in a country is an increasing function with regard to the market potential and the global factor productivity in this country, and decreasing with regard to fixed and production costs. For more simplicity, we present this relation as follows:

$$U_i(h) \equiv f\left(\underset{+}{\ln MP}, \underset{+}{\ln A_i}, \underset{-}{\ln w_i}, \underset{-}{\ln r_i}, \underset{-}{\ln t_i}, \underset{-}{\ln F_i}\right) \quad (8)$$

This specification represents the theoretical background for the empirical work that now follows.

4 Empirical work

4.1 A location choice model: the conditional logit

In this paper, we seek to study the factors determining a specific firm single choice between some unordered alternatives. Each subsidiary h chooses country i where it will locate. An unordered choice model particularly

well adapted to our question is the model developed by McFadden in 1974, i.e., the conditional logit (firm fixed-effects logit model).

The conditional logit is a discrete choice model based on profit maximization. In such a model, each firm compares the profits related to the different location alternatives, and selects, among the N alternatives, the location that will maximize its profit $\pi_i(h)$.

In our example, for a French firm h facing N alternatives, the profitability of choosing i can be written

$$\pi_i(h) = \beta' Z_i + \varepsilon_{hi} \quad (9)$$

with Z_i a vector of independent variables that vary between location alternatives, β the vector of estimated parameters and ε_{hi} a random error term, which corresponds to unobserved variables related to location i and affecting the choice of firm h .

Each French subsidiary h , faced with N choices, will select a location i if the expected profit $\pi_i(h)$ exceeds the expected profits $\pi_j(h)$, for all $j \in N$ alternative locations.

The model is made operational by a particular choice of distribution for the disturbances. Let Y_h be a random variable that indicates the choice made. Following McFadden, if and only if the N disturbances are independent and identically distributed with Weibull distribution, the probability that firm h chooses location i is given by:

$$Pr(Y_h = i) = \exp(\beta' Z_i) / \sum_{j=1}^N \exp(\beta' Z_j) \quad (10)$$

This intuitive formulation of the conditional logit model presents nevertheless some limits due to the assumption concerning the disturbances, which implies the property of Independence from Irrelevant Alternatives (IIA). According to the IIA property, the likelihood of making a choice is independent of the other alternatives. In practice, this assumption could be problematic. In order to mitigate this problem, in our econometric specification we introduce in addition to the explanatory variables, dummy variables representing the four different country groups forming our sample. With the assumption that the error terms are correlated only within country groups and not across groups, the dummy variables (defined in section 5) should capture this correlation and reduce the IIA problem.

4.2 Data description

Dependent variable

The location choice

Data concerning French firms location choice have been gathered from the Survey about French firms located abroad, conducted by the Directorate of Treasury and Economic Policy (DGTPE in French) in 2002. This Department collects from French Economic Missions abroad the census of units whose capital is owned by a French parent company by at least 10%. The survey includes among other information, the units date of creation and their French classification NAF93 code, data which is crucial for this analysis. The global database covers 21379 investment decisions realized in all economic sectors in 128 countries during the past century. In our study, we concentrate on the manufacturing industry, excluding the two-level NAF93 code DF "Coke, Petroleum Refining and Nuclear Industry" which corresponds to specific sectors whose location determinants are beyond the scope of this study. In the end, reduced to manufacturing sectors and limited to data availability for the explanatory variables, our empirical sample covers 1332 French investments in 55 countries from 1996 to 2002⁴.

Explanatory variables

As we have seen in our theoretical model, the profitability of a location for a firm depends on the market potential of the location and its total factor productivity, the firm's production costs, and the fixed costs that correspond to launching an activity. We have calculated market potential and total factor productivity growth, following commonly used methods described hereafter. Production and fixed costs are represented using several proxies that we present below.

The market potential

The market potential is a general concept regarding the impact of demand on firms location. Gross domestic product of the host country or its population are the most commonly used proxies for demand variables, but they are very partial. Indeed, measuring the local demand through these variables present the major inconvenience to not take into account the demands emanating from nearby countries and the facility or difficulty to reach them. In this study we exceed these limits by using the concept of market potential of a location. Here, market potential (MP) means "demand accessibility". This concept was introduced by Harris (1954) who proposed, as measure of the potential demand that a firm faces, the sum of economic sizes of surrounding markets weighted over distances: $MP_i = \sum_{j=1}^N \frac{GDP_j}{dist_{ij}}$. The concept was then validated by its deduction from the standard model of the new trade theory, such as presented in the equation (4) (Krugman, 1992; Head and Mayer, 2004). The presentation of the market potential à la Harris is very insufficient because of omission of the price index, G_j , which allows taking into account the effect of the competition. Besides, this simplification supposes that the simple distance includes all costs, while the literature on the border effects on trade between countries refutes this hypothesis by underlining the importance of obstacles bound to borders (McCallum, 1995; Head and Mayer, 2000). In this study analyzing the determinants of French firms location choice, it seems then essential to consider, besides the distance, additional trade costs induced by crossing borders and sharing or not a common language, while estimating the market potential of all possible destinations. Following Redding and

⁴List of countries in Appendix C.

Venables (2004) and Head and Mayer (2004), we build a measure of market potential that aggregates the local demand and the demands emanating from nearby markets, while taking into account the effect of demand's depreciation due to obstacles related to shipping goods in space and across borders. The estimation technique is presented in Appendix B.

Moreover, given that market potential values have been calculated using estimators of trade regression⁵, the standard-errors of equation (10) are biased as they include also trade equation's errors. We use the bootstrap technique in order to obtain correct standard-errors (limiting also measurement-error-induced bias resulting from using a calculation technique for total factor productivity variable). Standard-errors reported in clogit regressions table are thus based on 100 bootstrap replications (see Efron and Tibshirani, 1993, for more details).

Total factor productivity

While there is no available data on total factor productivity, we apply here a "Growth Accounting" calculation method, following the technique developed by Robert Solow (1956) to calculate the rate of technological progress (*TFP_growth*). According to this technique, the sources of output growth are: contribution of capital growth = $\alpha \frac{\Delta K}{K}$; contribution of labor growth = $\beta \frac{\Delta L}{L}$; and contribution of total factor productivity growth = $\frac{\Delta TFP}{TFP}$. To resume, the technological progress in our study equals output growth which is not explained by factors growth ($TFP_growth = \frac{\Delta PIB}{PIB} - \alpha \frac{\Delta K}{K} - \beta \frac{\Delta L}{L}$)⁶.

Factor costs

Following our theoretical model, the total production cost faced by a firm is composed by labor, capital and environmental costs.

Suitable variables to reflect the capital and the labor costs would have been respectively the capital interest rate and the wages in manufacturing sectors. On account of many missing values in the International Labor Office Laborsta database, related to international labor costs, and because of an expected high correlation between these costs and governance factors, we have finally decided to capture capital and labor costs through the countries' relative endowments in production factors. This is represented by the variable *KL*, which is the ratio K/L, with K the capital stock⁷ and L the total labor force.

The most complex cost to be represented is the environmental regulation (*ER*). Since a direct measure does not exist for this cost, we had recourse to different proxies that allowed us to compute, through the Z-score method⁸, a global and quite exhaustive Environmental Regulation index (ER index) for each country in our

⁵Cf Appendix B.

⁶The coefficients α and β are obtained by running regression: $\ln PIB = \alpha \ln K + \beta \ln L + \varepsilon$, for each of our country-groups, separately.

⁷The capital stock is calculated by using the following formula: $K_t = \text{gross fixed capital formation}_t + 0.95 \text{ capital stock}_{t-1}$. Due to data availability (particularly concerning transition countries), the initial stock is represented by the gross fixed capital formation in 1990.

⁸We first calculate for each variable and year, the distance between each country's value and the mean of the group expressed in standard-errors, following the formula: $z = (X_{jt} - \bar{X}_t) / \sigma_t$. We obtain thus values allowing to classify the countries below or above the mean. Then, we calculate the unweighted average of all variables' z-scores. Finally, we apply the *standard normal percentile* technique which gives the value 0 to the least average Z-score and 100 to the highest. Thus, the value 50 corresponds to the mean of the sample.

sample. This index integrates the following variables:

Multilateral Environmental Agreements (MEA ratified): this variable distinguishes countries having ratified several international environmental agreements. Thus, countries ratifying more MEAs prove their governments' concern about environmental protection. We retained the MEA's ratification year (and not the year of signature) since it is the ratification that imposes compliance to international environmental treaties.

ISO 14001 (ISO 14001 certifications/billion US\$ GDP): we have integrated for each country the number of ISO 14001 certifications normalized by the country's GDP. Even if ISO certification is a private and voluntary initiative, this variable manages to express a global state of mind prevailing in a country. For illustration, this variable is also used in the construction of the Environment Sustainability Index jointly initiated by the Yale Center for Environmental Law and Policy and Columbia University. On one hand, one can suppose that a higher number of 14001 ISO certified firms is the consequence of strict standards and controls imposed by the government. On the other hand, countries where this variable is the most important should be considered as countries where population is the most sensitive to environmental issues and thus the most exacting to environmental policy's severity.

International NGOs (INGOs' members/million of population): this variable represents the density of international non-governmental organizations with membership. As mentioned by Dasgupta et al. (1995) and Smarzynska and Wei (2004), international NGOs make local population sensitive to environmental problems, and also put pressure on governments to respect laws. Thus, a more important presence of international NGOs in a country would imply a more stringent environmental regulation.

Energy efficiency (GDP/unit of energy used): the interest of using such a quantitative variable is that it gives a real measure of the impact of the preceding variables⁹. This allows distinguishing countries that apply concrete environmental measures from the ones that adopt a "theoretical" environmental policy not really restrictive to firms.

Thus, grouping these variables in a single index allows us to encompass the general environmental regulation of countries according to different more precise environmental aspects. For instance, some countries can have a small number of ISO 14001 certified firms, but high energy efficiency. The use of one variable rather than the other in the regression would then give an incomplete vision of the local environmental regulation and result in wrong *ER* coefficients. As illustration, we can cite the work of Smarzynska and Wei (2004) who, by using several too precise environmental variables, taken separately in the estimations, did not manage to capture the general aspect of the environmental regulation, which may have prevented them from proving explicitly the pollution haven hypothesis. Moreover, introducing a unique proxy for environmental regulation stringency can lead to biased results. For example, if we only use the MEA variable as the environmental proxy without controlling for other aspects of the environmental policy climate, it could capture in this case the state's willingness to keep

⁹A more appropriate quantitative measure would have been the manufacturing sectors' CO₂ emissions normalized by the value added in the respective sectors. Due to a lack of consistent data that would have prevented us from including several countries, we have preferred to consider energy efficiency, which generates similar results (available upon request).

reliable relationships on international arena, which is usually favorable for FDI, rather than its direct concern to international, and much less domestic environmental compliance.

Table 1 presents correlations among the four variables and the ER index. On the one hand, some variables are well correlated, e.g., ISO 14001 and NGO variables (0.46), or MEA and Energy efficiency variables (0.39), while others, e.g. ISO 14001 and Energy efficiency, have a correlation coefficient close to zero (0.092). On the other hand, we can see in the first column that all variables are strongly enough positively correlated with the ER index. This lends support to our argument that these variables measure each one a different aspect of environmental stringency, and that taken together they represent environmental stringency in general.

| | ER index | MEA | ISO 14001 | NGO | Energy efficiency |
|-------------------|----------|-------|-----------|-------|-------------------|
| Index ER | 1.000 | | | | |
| MEA | 0.534 | 1.000 | | | |
| ISO 14001 | 0.496 | 0.267 | 1.000 | | |
| NGO | 0.633 | 0.247 | 0.456 | 1.000 | |
| Energy efficiency | 0.585 | 0.389 | 0.092 | 0.118 | 1.000 |

Table 1: Correlation of environmental index components

Governance factors

Since there is no variable representing precisely a firm’s fixed costs to establish itself in a location, we have used proxies of institutional quality in a location in order to capture these costs. Indeed, bad governance generates additional costs while launching an activity and creates a feeling of insecurity among investors. Factors related to country governance constitute then determining factors to firms’ location abroad, especially when we consider developing countries or transition economies, where governance problems are rather frequent. In our paper, we use as proxies three governance indicators developed by Kaufmann et al. (2005): the corruption level (*CORRUP*), the political stability (*POLSTAB*) and the government regulatory quality (*REGULQUAL*). *CORRUP* is the inverse of the original Kaufmann index which reflects the control of corruption in states (a higher value meaning a better governance outcome). Our corruption variable should then have a negative effect on location decisions as a result of greater corruption. At the opposite, we expect the attractiveness of a location to increase with political stability and absence of violence (*POLSTAB*), and the government ability to implement promoting regulations (*REGULQUAL*). Tables 4 and 5 in Appendix present descriptive statistics for the independent variables and information about the data.

5 Empirical results

Table 2 shows estimation coefficients concerning the impact of different factors on the location of French firms. Besides variables presented in the preceding section, our estimations include also dummy variables that we

have created by grouping countries in four homogenous clusters: *TrCEEC* for transition CEECs, *TrCIS* for transition countries of CIS, *EMERG* for emerging countries, and *OECD* for high-income OECD members. Variables are log-linearized, and *KL* and *ER* have been lagged one-year to avoid any possible endogeneity with the dependent variable.

| Variables | Conditional logit | | | | |
|-----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| | (1) | (2) | Model (3) | (4) | (5) |
| lnMP | 0.566*** (0.049) | 0.516*** (0.054) | 0.437*** (0.093) | 0.494*** (0.11) | 0.479*** (0.10) |
| TFP_growth | 1.667** (0.67) | 1.573** (0.66) | 2.330*** (0.84) | | 1.999** (0.80) |
| lnKL _{t-1} | -0.362*** (0.060) | -0.391*** (0.060) | -0.368*** (0.091) | -0.355*** (0.089) | -0.324*** (0.089) |
| lnER _{t-1} | -1.839*** (0.26) | -1.848*** (0.26) | -1.934*** (0.33) | -1.476*** (0.41) | -1.577*** (0.39) |
| lnCORRUP | -1.051*** (0.25) | -1.117*** (0.25) | -0.746* (0.45) | 0.0390 (0.48) | -0.213 (0.46) |
| lnPOLSTAB | 0.590** (0.22) | 0.777*** (0.21) | 0.534* (0.29) | 1.142** (0.46) | 0.960** (0.40) |
| lnREGULQUAL | 0.778*** (0.29) | 0.675** (0.29) | 1.418*** (0.42) | 1.054 (0.74) | 0.732 (0.66) |
| TrCEEC | -0.232* (0.14) | 0.0918 (0.18) | 0.186 (0.28) | 0.141 (0.30) | 0.219 (0.30) |
| TrCIS | -0.433** (0.21) | -0.220 (0.23) | -0.230 (0.40) | 0.154 (0.41) | 0.0162 (0.40) |
| EMERG | 0.515*** (0.12) | 0.808*** (0.16) | 0.866*** (0.27) | 1.138*** (0.34) | 1.123*** (0.32) |
| PTA | | 0.392*** (0.14) | 0.483** (0.24) | 0.702** (0.27) | 0.640** (0.25) |
| lnR&D | | | | 0.0713** (0.033) | 0.0667** (0.031) |
| Observations | 71097 | 71097 | 33065 | 21500 | 21486 |
| Pseudo R ² | 0.0356 | 0.0365 | 0.0318 | 0.0325 | 0.0339 |

Bootstrap standard-errors in parentheses. *** p<0.01 ** p<0.05 * p<0.1.

Table 2: Conditional logit estimates for the pooled sample of countries, coefficients

Our base model, estimation (1), shows results consistent with theory and our predictions for all the explanatory variables. First, concerning our core variable in this study, environmental regulation, it seems to be an important factor for French manufacturing firms' location decisions. The estimated coefficient of the environmental regulation index is negative and consistently significant at the 1% level, indicating that a more stringent environmental regulation deters French manufacturing investments. Everything else equal, all industries have interest to avoid additional costs induced by stricter environmental regulation, since there is generally no totally "clean" manufacturing industry. Second, market potential and total factor productivity growth appear to be important attractive factors for French direct investments abroad. Moreover, French firms are attracted by relatively labor-abundant countries, an increase in the ratio K/L having a negative and strongly significant effect on the location decision. Finally, host countries' governance also influences French firms' decision to settle or not in a country, since our three governance variables are significant and show the expected signs. Thus, politically stable countries attract French FDI, as well as a satisfactory regulatory quality, while a high level

of corruption discourages it. All our dummy variables representing country groups are significant, and indicate that between 1996 and 2002, French firms preferred to establish predominantly in emerging countries compared to OECD countries, but much less in transition economies.

In order to assert these main results and control for any potential effect on location of trade openness between host countries and France, we have incorporated to the base model the variable *PTA* (columns (2) to (5)). Variable *PTA* takes value 1 if a country is a EU-member or has contracted a preferential trade agreement with the European Union, and 0 otherwise. We observe in column (2) that *PTA* has the expected sign for the pollution haven hypothesis to work (under free trade, relocation of pollution-intensive production takes place from countries with stringent pollution regulation to countries with lax regulation countries) and it is significant at the 1% level, while other variables and especially the ER index maintain their sign and significance. Dummy variables TrCEEC and TrCIS are no more significant, perhaps because in this case French firms choose countries depending on the existence of a trade agreement rather than their belonging to a specific country group.

To highlight PHH reality, we focus in model (3) on the particular case of firms belonging to the most polluting manufacturing sectors, defined here as sectors the abatement costs of which exceed 0.5% of value added in the classification provided by Raspiller and Riedinger (2004). These firms should be more affected by a stricter environmental regulation. We first notice that nearly all the coefficients keep high significance, which attests to the robustness of our results. Nevertheless, parameters estimates reveal some particular characteristics specific to polluting firms. Regarding the environmental regulation index, the coefficient remains very significant at the 1% level, with a higher magnitude, indicating the higher sensitivity of most polluting firms to environmental regulation. This result emphasizes the pollution haven hypothesis previously proven in models (1) and (2). Market potential and the K/L ratio coefficients have lower magnitudes, while the variables for political stability and corruption remain significant at the 10% level only. At the opposite, total factor productivity and regulatory quality show a higher impact on FDI location compared to the model including all industries. We can deduce from these findings that firms in the most polluting sectors, that often are more capital intensive and require heavier investments, may be relatively more attracted by capital endowments, total factor productivity and regulatory quality than the less polluting ones. For the same reasons, these firms are also geographically less mobile, and thus may be slightly less sensitive to market potential. All dummies remain quite robust.

In order to verify the robustness of these first empirical results, we run two additional regressions. Firstly, we introduce in model (4) the share of research and development expenditure in GDP (*R&D*) as a proxy for the total factor productivity, instead of *TFP_growth*. Secondly, we include the two variables simultaneously in model (5), considering that *TFP_growth* represents the effect of technological progress while *R&D* captures the impact of a given level of total factor productivity. Moreover, since rich countries have usually the largest share of *R&D* expenditure in GDP, the use of *R&D* as a proxy for the level of total factor productivity allows us to control for some effects of the country's economic development that could be correlated with the stringency of its environmental regulation. The regression results confirm the robustness of model (3), except for *CORRUP*

and *REGULQUAL* which are no more statistically significant in both models. The variables *TFP_growth* and *R&D* are statistically significant at the 5% level and have the expected positive signs. As for the environmental regulation, even after an additional control for the country’s development level, it maintains a negative effect on FDI location, significant at the 1% level. Considering the high number of missing data on R&D, we concentrate further sensitivity analysis on model (3).

Country-group analysis

Based on the previous results, which provide evidence in favour of the pollution haven hypothesis, we extend our analysis in order to distinguish which countries are the most likely to constitute pollution havens. To this goal, we need to introduce interaction terms between the index of environmental regulation and the country-group dummies. However, as noted by Ai and Norton (2003), the interaction effect’s magnitude and significance should not be based on the interaction coefficient in non-linear models, e.g. the conditional logit, because its magnitude and significance may vary across the range of predicted values. Consequently, given that interaction terms cannot be correctly interpreted from our conditional logit model results, we first run estimation of model (3) through a logit model with adjusted standard-errors for intragroup correlation¹⁰. Since the results are consistent with those found with the conditional logit, we can then use the Norton’s et al. (2004) methodology recommended for computing the marginal effects of the interaction terms in logit models.

In Table 3, we present the marginal effects of the reference model (3) logit estimations with *ER* impact by country-group. Apart from the ER index, and except the corruption variable, which is not significant in models (7) and (9), our explanatory variables maintain their sign and significance across all models (6) to (9), as compared to model (3)¹¹. Concerning the environmental regulation, in models (6) to (9), where the country group in the interaction term varies, the sign of the coefficient of the *ER* variable indicates once again that a stricter environmental regulation deters foreign investments. Concerning the effect for different country groups, we observe that the interaction term is not significant for CEECs, emerging and OECD countries (models (6), (8) and (9) respectively), suggesting that the effect of environmental regulation for these countries is identical to the *ER* variable’s marginal effect, i.e., a negative and significant effect, and does not differ across these country groups. By comparison, model (7) for CIS countries, yields a significant interaction term, and results in a positive marginal effect of about 0.157, which indicates that a strengthened environmental regulation in this country group attracts French firms. These results suggest that the pollution haven hypothesis does not hold for CIS countries, while it is uniformly confirmed for the other country groups.

¹⁰ Estimation results available from the authors upon request.

¹¹ Comparing estimation coefficients, not marginal effects, of Tables 2 and 3. The estimation coefficients from Table 3 are available upon request.

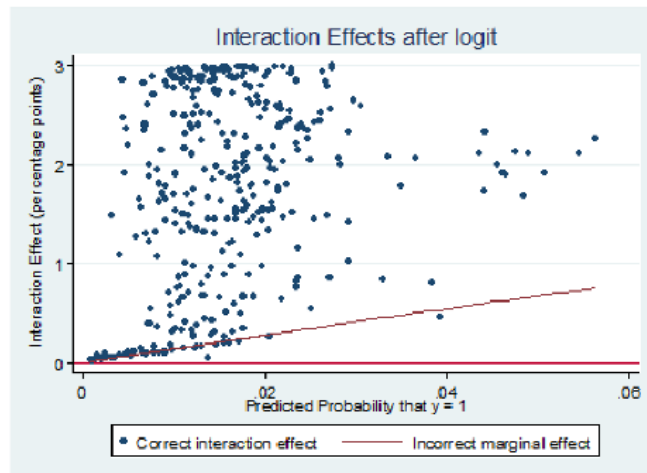
| Logit estimations (adjusted standard-errors for intragroup correlation) | | | | |
|---|----------------------|----------------------|----------------------|----------------------|
| Variables | Models | | | |
| | (6) | (7) | (8) | (9) |
| lnMP | 0.006*** (0.001) | 0.006*** (0.001) | 0.006*** (0.001) | 0.006*** (0.001) |
| TFP_growth | 0.024*** (0.008) | 0.023*** (0.007) | 0.024*** (0.007) | .024*** (0.007) |
| lnKL _{t-1} | -0.006*** (0.001) | -0.005*** (0.001) | -0.006*** (0.001) | -0.006*** (0.001) |
| lnER _{t-1} | -0.027*** (0.005) | -0.031*** (0.005) | -0.023*** (0.005) | -0.020** (0.008) |
| lnER*TrCEECC | 0.004 (0.015) | | | |
| lnER*TrCIS | | 0.188*** (0.029) | | |
| lnER*EMERG | | | -0.017 (0.012) | |
| lnER*OECD | | | | -0.011 (0.011) |
| lnCORRUP | -0.011* (0.006) | -0.007 (0.006) | -0.011* (0.006) | -0.010 (0.006) |
| lnPOLSTAB | 0.008* (0.004) | 0.012*** (0.004) | 0.007* (0.004) | 0.008* (0.004) |
| lnREGULQUAL | 0.021*** (0.006) | 0.022*** (0.006) | 0.022*** (0.006) | 0.021*** (0.006) |
| PTA | 0.007** (0.003) | 0.008** (0.003) | 0.007** (0.003) | 0.007** (0.003) |
| Country-group dummies | Yes | Yes | Yes | Yes |
| Observations | 38586 | 38586 | 38586 | 38586 |
| Pseudo R ² | 0.0249 | 0.03 | 0.0252 | 0.0250 |

Adjusted standard-errors in parentheses. *** p<0.01 ** p<0.05 * p<0.1.

Table 3: Interaction terms in logit estimations, marginal effects

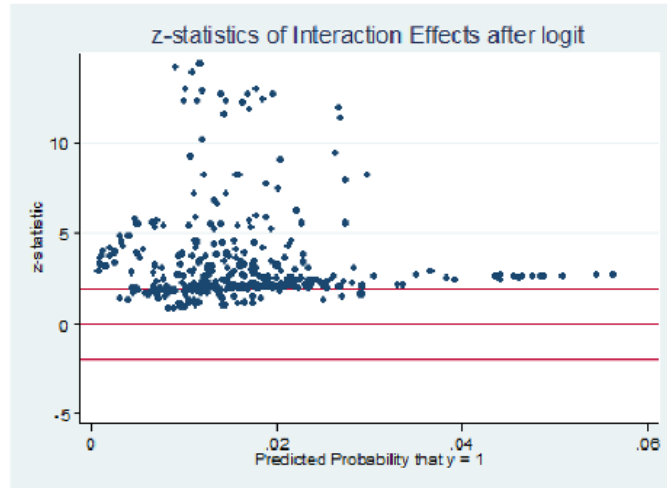
However, as previously mentioned, the interaction effect's magnitude and significance may vary across the range of predicted values, and this early conclusion can be misleading. The methodology recommended by Norton et al. (2004) allows us to visualize the correct interaction effect through two figures.

We present below the interaction effect for CIS countries.



Positive effect for CIS countries.

We observe that the interaction effect is positive across all observations, and that it offsets the negative *ER* marginal effect of -0.031 in model (7) since it takes values superior to 0.031 in most cases. In addition, Figure 2 shows that it is significant across almost the entire range of predicted probabilities of choosing a country (X-axis), since most of observations have a z-statistic superior to the 1.96 value (represented by the upper horizontal line on the Y-axis) corresponding to the 5% level of significance.

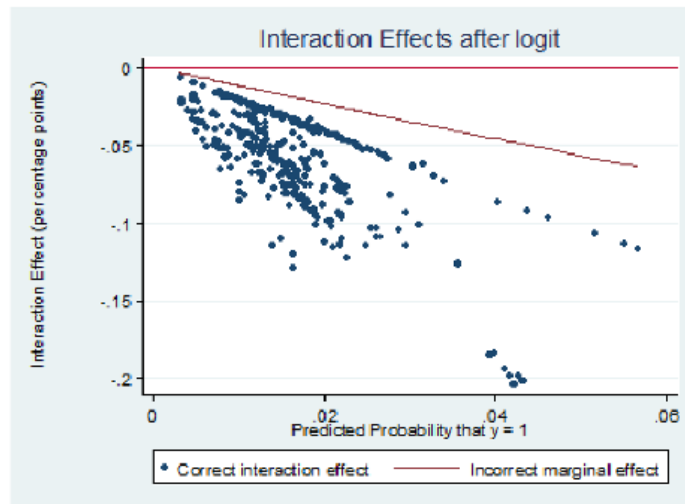


Significance of the interaction coefficient, CIS countries.

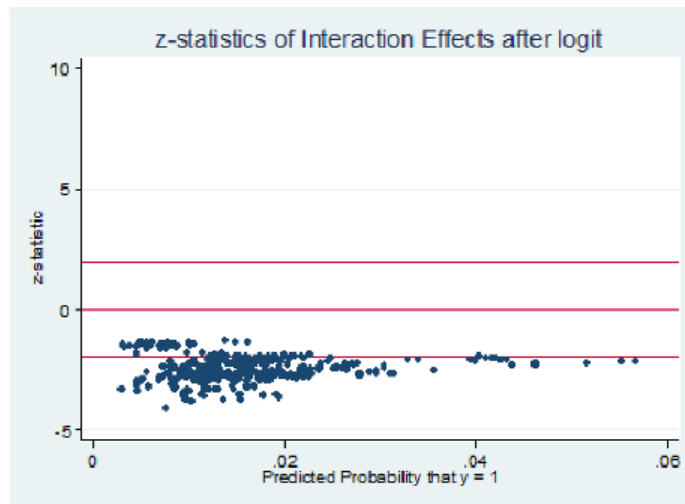
Thus, the interpretation of the interaction effect compared to the *ER* marginal effect of -0.031 mentions that a stricter regulation in CIS countries generally attracts foreign investments, with the pollution haven hypothesis nevertheless validated for a minor part of CIS observations. This predominant opposite effect for CIS countries could indicate some reluctance of French firms to locate in countries where the environmental regulation is too lenient.

When we apply the above methodology to other country groups, we find that for CEECs and OECD countries the marginal interaction effect, although predominantly positive, is not significant, which means that the impact of environmental regulation for these country groups is the same as for the base country group in models (6) and (9) respectively, i.e., significantly negative.

The most obvious pollution haven effect appears for emerging countries. Thus, as we can observe in Figures 3 and 4, the interaction effect is negative for all observations related to emerging countries, and significant for most of them.



Negative effect for emerging countries.



Significance of the interaction coefficient, emerging countries.

We can thus say that the negative effect of environmental regulation is the strongest for emerging countries, compared to the base country group.

6 Conclusion

In this study we have tested the pollution haven hypothesis through an analysis of the impact of environmental regulation on French manufacturing firms' location choice. Using firm-level data concerning French firms' locations in the world, we have first tested this hypothesis for a pooled sample, and then tested it making a distinction between four country groups: transition CEECs, transition countries of CIS, emerging countries, and high-income OECD countries. By applying a geographic economy model, which has the advantage of considering a complete set of FDI determinants like market potential, production factors and governance quality, and by developing a complex index encompassing the different aspects of environmental regulation, we have succeeded

in expressing the stringency of environmental regulation in a satisfying way and in revealing thus the existence of pollution havens.

Empirical results of the base model show that in presence of heterogeneous countries, French manufacturing industries locate in countries with more lenient regulations, thus confirming the essential role played by environmental regulation in determining firms' location. Moreover, this effect is reinforced for the most polluting firms.

In order to test the robustness of our empirical results and considering only the firms belonging to the most polluting industries, we identified the country groups that are the most likely to constitute pollution havens. Estimations including interaction terms between environmental regulation and country groups validated the pollution haven hypothesis for CEECs, emerging and OECD countries included in our sample. On the contrary, concerning CIS countries, a more stringent regulation seems rather to attract investments. Finally, we show that the pollution hypothesis holds the strongest for emerging economies.

As for the policy implications, we conclude that the approval or rejection of the pollution haven hypothesis is not sufficient to respond to fears related to the impact on firm location of heterogeneous environmental regulations between economies. Indeed, although our study has validated the pollution haven hypothesis for a pooled sample of countries, it highlights that too large a gap in the stringency of environmental regulation between countries deters foreign investments from the less regulated ones, usually poorer countries. Consequently, a too lenient environmental policy in the less developed countries could be to the detriment of their technological modernization that might result from FDI externalities, and thus prevent potential host countries from improving their environmental quality. Research examining to which extent pollution havens imply a real threat to the environment, or at the opposite could be beneficial to it thanks to technological improvements for example, would be of a great interest.

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Appendices

A Theory - Deriving the demand function for varieties in a geographic economy model

Following the general description of the model in section 3, consumers spend a part $0 < \mu < 1$ of their income E on the purchase of the composite good M . Their preferences in a country j are described by the following utility function:

$$U_j = C_{Aj}^{1-\mu} C_{Mj}^\mu \quad (\text{A. 1})$$

where C_A is the consumption of the traditional good A and C_M is the consumption of the manufacturing composite good M , for which consumers have CES sub-utility functions. CES preferences are at the heart of the Dixit-Stiglitz monopolistic competition model (Dixit and Stiglitz, 1977). In our case, they are expressed in terms of a continuum of varieties:

$$C_{Mj} = \left(\sum_i^N \int_{n_i} q_{ij}(v)^{1-1/\sigma} dv \right)^{\frac{1}{1-1/\sigma}} \quad (\text{A. 2})$$

with $0 < \mu < 1 < \sigma$; n_i is the mass of varieties produced in a country i , $i \in N$; N - the number of countries in the world; $q_{ij}(v)$ - is the consumption of the v^{th} variety in the country j and $\sigma > 1$ is the constant elasticity of substitution. The corresponding indirect utility function can be written as E/P_M where the price index P_M in the country j is

$$P_{Mj} \equiv P_{Aj}^{1-\mu} \left(\sum_i^N \int_{n_i} p_{ij}(v)^{1-\sigma} dv \right)^{\frac{\mu}{1-\sigma}} \quad (\text{A. 3})$$

$P_A = 1$ is the price in the normalized sector A and $p_{ij}(v)$ is the consumption price of the v^{th} industrial variety produced in country i and sent in country j . P_M is also called a "perfect" price index since it translates expenditure, μE , into utility.

Under the budget constraint:

$$L_j = \left(\sum_i^N \int_{n_i} q_{ij}(v)^{1-1/\sigma} dv \right)^{\frac{1}{1-1/\sigma}} - \lambda \left[\sum_i^N \int_{n_i} p_{ij}(v) q_{ij}(v) dv - \mu E_j \right] \quad (\text{A. 4})$$

we obtain the following first order condition of CES sub-utility maximization:

$$[q_{ij}(h)]^{-1/\sigma} \left(\sum_i^N \int_{n_i} q_{ij}(v)^{1-1/\sigma} dv \right)^{\frac{1}{1-1/\sigma}-1} = \lambda p_{ij}(h) \quad (\text{A. 5})$$

with $q_{ij}(h)$ and $p_{ij}(h)$ - the consumption and price, respectively, of an alternative/specific variety h , $h \in [1, n_i]$.

Next, we multiply both sides by $q_{ij}(h)$, take the integral of varieties and sum across the entire mass of varieties produced in all N countries. Using the budget constraint, $\mu E_j = \sum_i \int_{n_i} p_{ij}(v) q_{ij}(v) dv$, we obtain an expression for the Lagrangian multiplier, $\lambda = \left(\sum_i \int_{n_i} q_{ij}(v)^{1-1/\sigma} dv \right)^{\frac{1}{1-1/\sigma}} (\mu E_j)^{-1}$. Isolating $q_{ij}(h)$ on the left-hand side, we have:

$$[q_{ij}(h)]^{1/\sigma} = \left(\sum_i \int_{n_i} q_{ij}(v)^{1-1/\sigma} dv \right)^{\frac{1}{1-1/\sigma}-1} [\lambda p_{ij}(h)]^{-1} \quad (\text{A. 6})$$

Substituting the expression for λ , we obtain:

$$[q_{ij}(h)]^{1/\sigma} = \frac{\mu E_j \left(\sum_i \int_{n_i} q_{ij}(v)^{1-1/\sigma} dv \right)^{\frac{1}{1-1/\sigma}-1}}{p_{ij}(h) \left(\sum_i \int_{n_i} q_{ij}(v)^{1-1/\sigma} dv \right)^{\frac{1}{1-1/\sigma}}} \quad (\text{A. 7})$$

Applying some transformations to equation (A. 7), we obtain the direct and inverse demand curves, respectively.

These are:

$$q_{ij}(h) = \frac{[p_{ij}(h)]^{-\sigma}}{\sum_i \int_{n_i} [p_{ij}(v)]^{1-\sigma} dv} \mu E_j; \quad p_{ij}(h) = \frac{[q_{ij}(h)]^{-1/\sigma}}{\sum_i \int_{n_i} [q_{ij}(v)]^{1-1/\sigma} dv} \mu E_j; \quad (\text{A. 8})$$

B Market Potential estimation

Krugman's market potential has the advantage of being deduced strictly from theory:

$$MP_i = \sum_j \left[\frac{\phi_{ij} \mu E_j}{G_j} \right] \quad (\text{B. 1})$$

where $G_j = \sum_i \int_{n_i} [c_i(v) \tau_{ij}]^{1-\sigma} dv$. Nevertheless, compared to the form proposed by Harris (1954), its calculation needs estimators for the unknown φ_{ij} and G_j parameters. In this study we apply the same strategy as Head and Mayer (2004) - we estimate these parameters using information about international trade flows. The aggregate value of exports of country i towards country j , denoted here by X_{ij} , results by multiplying the mass of varieties produced in country i and sent to country j by each variety's export price (including trade costs):

$$X_{ij} = \int_{n_i} p_{ij}(v) q_{ij}(v) dv = \int_{n_i} \frac{c_i(v)^{1-\sigma} \phi_{ij} \mu E_j}{G_j} dv \quad (\text{B. 2})$$

All variables are specified in section 3.

By grouping the terms according to the indexes and then transforming them in logarithm, we obtain:

$$\ln X_{ij} = \ln \left(\int_{n_i} c_i(v)^{1-\sigma} dv \right) + \ln (\mu E_j / G_j) + \ln \varphi_{ij} \quad (\text{B. 3})$$

Following Redding and Venables (2004), we estimate the first two terms by using exporter and importer fixed effects, denoted here EX_i and IM_j , respectively. The bilateral access to the market (φ_{ij}) is considered in a similar way as in Head and Mayer (2004) to be a function of distance (d_{ij}), contiguity ($B_{ij} = 1$ if countries i and j share a common border and 0 otherwise), common language ($L_{ij} = 1$ if i and j share a language and 0 otherwise) and an error term, ϵ_{ij} . The trade equation to be estimated is then:

$$\ln X_{ij} = EX_i + IM_j - \delta \ln d_{ij} + \beta B_{ij} + \lambda L_{ij} + \epsilon_{ij} \quad (\text{B. 4})$$

This equation is regressed on the bilateral trade flows of 168 countries over the period 1990-2000 (Feenstra's database on world trade flows, NBER) and 79 countries over the period 2001-2004 (Chelem database). The variables necessary to the calculation of ϕ_{ij} are taken from CEPII's Distances database.

Using the specifications $\hat{\phi}_{ij} = d_{ij}^{-\hat{\delta}} \exp \left(\hat{\beta} B_{ij} + \hat{\lambda} L_{ij} \right)$ and $\mu E_j / G_j = \exp (IM_j)$, we calculate the market potential.

C List of countries in the sample

Transition countries, CEECs: Bulgaria - Czech Republic - Estonia - Hungary - Latvia - Lithuania - Poland - Romania - Slovak Republic - Slovenia.

Transition countries, CIS: Azerbaijan - Georgia - Kazakhstan - Uzbekistan - Ukraine - Russia.

Emerging economies: Argentina - Brazil - Chile - China - Colombia - Egypt - India - Indonesia - Iran - Israel - Malaysia - Mexico - Morocco - Pakistan - Peru - the Philippines - Republic of Korea - Singapore - South Africa - Thailand - Turkey - Venezuela.

High-income OECD countries: Australia - Austria - Canada - Denmark - Finland - Germany - Greece - Italy - Ireland - New Zealand - Netherlands - Norway - Portugal - Spain - Sweden - Switzerland - United Kingdom.

D Summary statistics

| Variable | Mean | Standard deviation | Min | Max |
|-------------------------|--------|--------------------|-------|--------|
| <i>MP</i> | 12.97 | 13.32 | 2.12 | 68.46 |
| <i>TFP_growth</i> | -0.094 | 0.08 | -0.84 | 0.10 |
| <i>KL_{t-1}</i> | 33874 | 37764 | 1139 | 204495 |
| <i>ER_{t-1}</i> | 54.68 | 8.90 | 34.04 | 84.94 |
| <i>CORRUP</i> | 30.95 | 9.91 | 17.91 | 54.20 |
| <i>POLSTAB</i> | 3.401 | 0.88 | 1.05 | 4.76 |
| <i>REGULQUAL</i> | 4.60 | 0.85 | 2.18 | 6.31 |
| <i>PTA</i> | 0.35 | 0.48 | 0 | 1 |
| <i>RD</i> | 8.03 | 60.31 | 0.08 | 569.10 |

Table 4: Descriptive statistics, model (3)

E Data summary

| Variable | Définition | Sources |
|--------------------|---|---|
| Dependent variable | Location choice. Market potential. | Subsidiaries-Survey 2002, Economic Policy and Treasury General Direction (DGTPE). Data on international trade: R.Feenstra and R.Lipsey, NBER 1990-2000. Chelem, CEPII, 2000-2004. Geographical data: CEPII. World development Indicators, World Bank. |
| <i>KL</i> | Countries relative endowments in production factors (capital versus labor). | World development Indicators, World Bank. |
| <i>ER</i> | Environmental regulation index. | Authors calculation. |
| | Multilateral Environmental Agreements ISO 14001 | Earthtrends, World Resources Institute. International Organization for Standardization. Center for the Study of Global Governance. World Bank. |
| | International NGOs | |
| | Energy efficiency (GDP/unit of energy used) Corruption. | Governance Indicators 1996-2004, D. Kaufmann, A. Kray and M. Mastruzzi. Governance Indicators 1996-2004, D. Kaufmann, A. Kray and M. Mastruzzi. Governance Indicators 1996-2004, D. Kaufmann, A. Kray and M. Mastruzzi. |
| <i>CORRUP</i> | | Preferential trade agreements database. Multiple sources. |
| <i>POLSTAB</i> | Political Stability. | Multiple sources. |
| <i>REGULQUAL</i> | Regulations improving firms' general environment. Preferential trade agreements with EU. | Morgan Stanley emerging Markets index. World Bank Country Classification |
| <i>PTA</i> | Central and Eastern European Countries. Commonwealth of Independent States. | |
| <i>T*CEEC</i> | Emerging countries. | |
| <i>T*CIS</i> | | |
| <i>EMERG</i> | | |
| <i>OECD</i> | High-income OECD countries. | |

Table 5: Data definitions and sources.

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